# Probability in High Dimensions

Solutions to Exercises, Week 2

July 25, 2020

### 2 Variance bounds and Poincaré inequalities

#### 2.1Tensorization and bounded differences

**Problem 2.1** (Banach-valued sums). Let  $X_1, \ldots, X_n$  be independent random variables with values in a Banach space  $(B, \|\cdot\|_B)$ . Suppose these random variables are bounded in the sense that  $\|X_i\|_B \leq C$  a.s. for every *i*. Show that

$$\operatorname{Var}\left(\left\|\frac{1}{n}\sum_{k=1}^{n}X_{k}\right\|_{B}\right) \leq \frac{C^{2}}{n}.$$

Solution. Alternative 1:

Let  $z_1, z_2, a$  be elements of B with norm bounded by C and  $n \in \mathbb{N}$ . Then

$$\begin{aligned} \left\|\frac{z_1}{n} + a\right\|_B &\leq \left\|\frac{z_1}{n} - \frac{z_2}{n}\right\|_B + \left\|\frac{z_2}{n} + a\right\|_B \\ &\leq \frac{2C}{n} + \left\|\frac{z_2}{n} + a\right\|_B \end{aligned}$$

This implies

$$\sup_{\|z_1\|_B \le C} \left\| \frac{z_1}{n} + a \right\|_B \le \frac{2C}{n} + \left\| \frac{z_2}{n} + a \right\|_B$$

which in turn yields

$$\sup_{\|z_1\|_B \le C} \left\| \frac{z_1}{n} + a \right\|_B \le \frac{2C}{n} + \inf_{\|z_2\|_B \le C} \left\| \frac{z_2}{n} + a \right\|_B.$$

Set  $a = \frac{1}{n} \sum_{k \neq i} x_k$  and  $f(x_1, \dots, x_n) = \left\| \frac{1}{n} \sum_{i=1}^n x_i \right\|_B$  then

$$D_i f(x) = \sup_{\|z\|_B \le C} \left\| \frac{z}{n} + \frac{1}{n} \sum_{k \ne i} x_k \right\|_B - \inf_{\|z\|_B \le C} \left\| \frac{z}{n} + \frac{1}{n} \sum_{k \ne i} x_k \right\|_B$$
$$\leq \frac{2C}{n}.$$

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By Corollary 2.4 (Bounded Differences) we obtain

$$\operatorname{Var}(f(X_1, \dots, X_n)) \le \frac{1}{4} \mathbf{E} \left[ \sum_{i=1}^n (D_i f)^2 \right] \le \frac{1}{4} \left[ \sum_{i=1}^n \frac{4C^2}{n^2} \right] = \frac{C^2}{n}.$$

### Alternative 2:

In the spirit of an earlier version of this solution it is also possible to find the same bound by the reverse triangle inequality

$$D_i f(x) = \sup_{\|z_1\|_B \le C} \left\| \frac{z}{n} + a \right\|_B + \sup_{\|z_2\|_B \le C} - \left\| \frac{z}{n} + a \right\|_B$$
  
=  $\sup \left\{ \left\| \frac{z_1}{n} + a \right\|_B - \left\| \frac{z_2}{n} + a \right\|_B : \|z_1\|_B \le C, \|z_2\|_B \le C \right\}$   
 $\le \frac{1}{n} \sup \left\{ \|z_1 - z_2\|_B : \|z_1\|_B \le C, \|z_2\|_B \le C \right\} \le \frac{2C}{n}.$ 

Alternative 3:

$$\begin{aligned} \|\frac{1}{n}\sum_{i=1}^{n}X_{i}\|_{B} &= \sup_{t\in B^{\circ}}\left\langle t,\frac{1}{n}\sum_{i=1}^{n}X_{i}\right\rangle \\ D_{i}^{-} &= \sup_{t\in B^{\circ}}\left\langle t,\frac{1}{n}\sum_{i=1}^{n}X_{i}\right\rangle - \inf_{\|z_{i}\|\leq C}\sup_{t\in B^{\circ}}\left\langle t,\frac{1}{n}\sum_{j\neq i}X_{i} + \frac{z_{i}}{n}\right\rangle \\ &= \sup_{\|z_{i}\|\leq C}\left[\sup_{t\in B^{\circ}}\left\langle t,\frac{1}{n}\sum_{i=1}^{n}X_{i}\right\rangle - \sup_{t\in B^{\circ}}\left\langle t,\frac{1}{n}\sum_{j\neq i}X_{i} + \frac{z_{i}}{n}\right\rangle\right] \end{aligned}$$

and since  $\sup f - \sup g \leq \sup[f - g]$ 

$$= \sup_{\|z_i\| \le C} \left[ \sup_{t \in B^{\circ}} \left\{ \left\langle t, \frac{1}{n} \sum_{i=1}^{n} X_i \right\rangle - \left\langle t, \frac{1}{n} \sum_{j \ne i} X_i + \frac{z_i}{n} \right\rangle \right\} \right]$$
$$= \sup_{\|z_i\| \le C} \left[ \sup_{t \in B^{\circ}} \left\langle t, \frac{z_i}{n} \right\rangle \right]$$
$$\le \frac{C}{n} \sup_{\|w\| \le 1} \left[ \sup_{t \in B^{\circ}} \left\langle t, w \right\rangle \right]$$

and since  $\sup_{t \in B^{\circ}} \langle t, w \rangle = ||w||_B$ 

$$= \frac{C}{n} \sup_{\|w\| \le 1} [\|w\|_B] \le \frac{C}{n}$$

**Problem 2.2** (Rademacher process). Let  $\varepsilon_1, \ldots, \varepsilon_n$  be independent symmetric random Bernoulli random variables  $\mathbf{P}[\varepsilon_i = \pm 1] = \frac{1}{2}$  (also called Rademacher variable), let  $T \subseteq \mathbb{R}^n$ . The following identity is completely trivial:

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$$\sup_{t\in T} \operatorname{Var}\left[\sum_{k=1}^{n} \varepsilon_k t_k\right] = \sup_{t\in T} \sum_{k=1}^{n} t_k^2.$$

Prove the following nontrivial fact:

$$\operatorname{Var}\left[\sup_{t\in T}\sum_{k=1}^{n}\varepsilon_{k}t_{k}\right] \leq 4\sup_{t\in T}\sum_{k=1}^{n}t_{k}^{2}.$$

Thus taking the supremum inside the variance costs at most a constant factor.

Solution. Let  $\langle x, y \rangle := \sum_{k=1}^{n} x_k y_k$  denote the inner product between vectors x and y in  $\mathbb{R}^n$ . Define the function  $f : \varepsilon \in \{-1, 1\}^n \to f(\varepsilon) := \sup_{t \in T} \langle \varepsilon, t \rangle$ . If we assume that for any  $\varepsilon \in \{-1, 1\}^n$  the supremum in  $\sup_{t \in T} \langle \varepsilon, t \rangle$  is attained (this would be the case if the set T is compact, by the extreme value theorem), then we can let  $t(\varepsilon) \in \arg \max_{t \in T} \langle \varepsilon, t \rangle$  denote any of its optimizers. Then, for any  $\varepsilon \in \{-1, 1\}^n$  and any  $i \in \{1, \ldots, n\}$ , if we let  $z(\varepsilon) \in \arg \min_{z \in \{-1, 1\}} f(\varepsilon_1, \ldots, \varepsilon_{i-1}, z, \varepsilon_{i+1}, \ldots, \varepsilon_n)$  we have

$$\begin{split} D_i^- f(\varepsilon) &:= f(\varepsilon) - \min_{z \in \{-1,1\}} f(\varepsilon_1, \dots, \varepsilon_{i-1}, z, \varepsilon_{i+1}, \dots, \varepsilon_n) \\ &= f(\varepsilon) - f(\varepsilon_1, \dots, \varepsilon_{i-1}, z(\varepsilon), \varepsilon_{i+1}, \dots, \varepsilon_n) \\ &= \sup_{t \in T} \langle \varepsilon, t \rangle - \sup_{t \in T} \langle (\varepsilon_1, \dots, \varepsilon_{i-1}, z(\varepsilon), \varepsilon_{i+1}, \dots, \varepsilon_n), t \rangle \\ &\leq \langle \varepsilon, t(\varepsilon) \rangle - \langle (\varepsilon_1, \dots, \varepsilon_{i-1}, z(\varepsilon), \varepsilon_{i+1}, \dots, \varepsilon_n), t(\varepsilon) \rangle \\ &= \langle (0, \dots, 0, \varepsilon_i - z(\varepsilon), 0, \dots, 0), t(\varepsilon) \rangle \\ &\leq 2 |t(\varepsilon)_i|. \end{split}$$

The proof follows by applying Corollary 2.4 upon noticing that since  $t(\varepsilon) \in T$  we have

$$\sum_{i=1}^n (D_i^-f(\varepsilon))^2 \leq 4\sum_{i=1}^n t(\varepsilon)_i^2 \leq 4\sup_{t\in T}\sum_{i=1}^n t_i^2.$$

More generally if  $\mathcal{T}$  is bounded then by continuity we can replace it by its closure which is compact. Finally, if  $\mathcal{T}$  is unbounded there must be a basis vector  $\mathbf{e}_i$  such that  $\sup_{t\in\mathcal{T}}\langle \mathbf{e}_i,t\rangle = \infty$ , and since any basis vectors  $\mathbf{e}_i$  can be written as a linear combination of two elements of  $\{\pm 1\}^n$ , we can always find an element  $\varepsilon \in \{\pm 1\}^n$  such that  $\sup_{t\in\mathcal{T}}\langle \varepsilon,t\rangle = \infty$ . In this case, with  $\epsilon$  uniform in  $\{\pm 1\}^n$ , we have  $\sup_{t\in\mathcal{T}}\langle \epsilon,t\rangle = \infty$  with positive probability and the inequality holds trivially.

Problem 2.3 (Bin packing).

Solution. (a)

$$D_i f(X_1, \dots, X_n) = \sup_{z_i} f(X_1, \dots, X_{i-1}, z_i, X_{i+1}, \dots, X_n) - \inf_{z_i} f(X_1, \dots, X_{i-1}, z_i, X_{i+1}, \dots, X_n)$$

The sup will be at worst  $B_n + 1$  and the inf at worst  $B_n - 1$ . But, suppose that  $\inf_{z_i} = B_n - 1$ . This means that if we reduce the size of  $X_i$  to 0, then we gain a bin. Therefore, there is an optimal allocation with  $B_n$  bin in which  $X_i$  is a single occupant of a bin. Thus in this case, since  $X_i \leq 1$  we can increase its  $D_i \leq 1$ . (b) Trivial since the total size of the packages  $\sum X_i$  must fit in the total space of the bins which is  $B_n \times 1$ . Thus  $n\mathbb{E}[X_1] \leq \mathbb{E}[B_n]$ .

**Problem 2.4** (Order Statistics and spacings). Let  $X_1, ..., X_n$  be independent random variables, and denote by  $X_{(1)} \ge ... \ge X_{(n)}$  their decreasing rearrangement. Show that

$$\operatorname{Var}\left[X_{(k)}\right] \le k \mathbb{E}\left[\left(X_{(k)} - X_{(k+1)}\right)^2\right] \tag{1}$$

for  $1 \leq k \leq n/2$  and that

$$\operatorname{Var}\left[X_{(k)}\right] \le (n-k+1) \mathbb{E}\left[(X_{(k-1)} - X_{(k)})^2\right]$$
(2)

for  $n/2 < k \leq n$ .

Solution. To prove first (1), we note that for  $f(X) = X_{(k)}$ 

 $D_i^- f(X) = X_{(k)} - \inf_Z f(X_{1:i-1}, Z, X_{i+1:n})$ 

where

$$\inf_{Z} f(X_{1:i-1}, Z, X_{i+1:n}) = \begin{cases} X_{(k)} & \text{if } X_i < X_{(k)}, \\ X_{(k+1)} & \text{if } X_i \ge X_{(k)}. \end{cases}$$

Hence it follows by Corollary 2.4. that we have

$$\operatorname{Var}\left[X_{(k)}\right] \leq \mathbb{E}\left[\sum_{i=1}^{n} \left(D_{i}^{-}f\left(X\right)\right)^{2}\right]$$
$$\leq \mathbb{E}\left[\sum_{i=1}^{n} \mathbb{I}\left(X_{i} \geq X_{(k)}\right) \left(X_{(k)} - X_{(k+1)}\right)^{2}\right]$$
$$\leq k\mathbb{E}\left[\left(X_{(k)} - X_{(k+1)}\right)^{2}\right]$$

where the last inequality follows from the definition of  $X_{(k)}$ . One can verify that having ties is not an issue. This establishes (1) for  $1 \le k \le n-1$  but this inequality is not tight for large values of k. To establish (2), one uses the increasing rearrangement  $X_{(1)} \le ... \le X_{(n)}$  and performs similar calculations.

**Problem 2.5** (Convex Poincaré inequalities). Let  $X_1, \ldots, X_n$  be independent random variables taking values in [a, b]. The bounded difference inequalities of Corollary 2.4 estimate the variance  $\operatorname{Var}[f(X_1, \ldots, X_n)]$  in terms of discrete derivatives  $D_i f$  or  $D_i^- f$  of the function f. The goal of this problem is to show that if the function f is convex, then one can obtain a similar bound in terms of the ordinary notion of derivative  $\nabla_i f(x) = \frac{\partial f(x)}{\partial x_i}$  in  $\mathbb{R}^n$ .

a. Show that if  $g : \mathbb{R} \to \mathbb{R}$  is convex, then

$$g(y) - g(x) \ge g'(x)(y - x)$$
 for all  $x, y \in \mathbb{R}$ .

b. Show using part a. and Corollary 2.4 that if  $f : \mathbb{R}^n \to \mathbb{R}$  is convex, then

$$\operatorname{Var}[f(X_1, \dots, X_n)] \le (b-a)^2 \mathbf{E}[\|\nabla f(X_1, \dots, X_n)\|^2].$$

- c. Conclude that if f is convex and L-Lipschitz, i.e.,  $|f(x) f(y)| \leq L ||x y||$  for all  $x, y \in [a, b]^n$ , then  $\operatorname{Var}[f(X_1, \ldots, X_n)] \leq L^2(b-a)^2$ .
- Solution. a. This can be proven directly from the definition of convexity and the derivative: for any  $t \in [0, 1]$ and  $x, y \in \mathbb{R}$ ,

$$(1-t)g(x) + tg(y) \ge g(x+t(y-x))$$
$$g(y) - g(x) \ge \frac{g(x+t(y-x)) - g(x)}{t}$$

Taking the limit as  $t \to 0$  completes the proof.

b. Consider a function  $g_i : [a, b] \to [a, b]$ , defined as

$$g_i(z) \equiv f(x_1, \ldots, x_{i-1}, z, x_{i+1}, \ldots, x_n),$$

for some fixed  $\{x_1, \ldots, x_{i-1}, x_{i+1}, \ldots, x_n\} \in [a, b]^{n-1}$  and  $i \in \{1, \ldots, n\}$ . This function is convex since f is convex in all of its arguments. Thus, by part a.,

$$g_i(z) \ge g_i(x_i) + g'_i(x_i)(z - x_i)$$

for all  $z, x_i \in [a, b]$ . In other words, for all  $\{x_1, \ldots, x_n, z\} \in [a, b]^{n+1}$ ,

$$f(x_1, \ldots, x_{i-1}, z, x_{i+1}, \ldots, x_n) \ge f(x) + \nabla_i f(x)(z - x_i),$$

where we write  $x \equiv \{x_1, \ldots, x_n\}$ . Thus, we can bound  $D_i^- f(x)$ , for all  $i \in \{1, \ldots, n\}$ , as

$$D_i^- f(x) = f(x) - \inf_{a \le z \le b} f(x_1, \dots, x_{i-1}, z, x_{i+1}, \dots, x_n)$$
  
$$\leq f(x) - \left[ f(x) + \inf_{a \le z \le b} \nabla_i f(x)(z - x_i) \right]$$
  
$$= -\inf_{a \le z \le b} \nabla_i f(x)(z - x_i).$$

Finally, invoking Corollary 2.4, we have

$$\operatorname{Var}[f(X_1, \dots, X_n)] \leq \mathbf{E} \left[ \sum_{i=1}^n \left( D_i^- f(X_1, \dots, X_n) \right)^2 \right]$$
$$\leq \mathbf{E} \left[ \sum_{i=1}^n \left( \inf_{a \leq z \leq b} \nabla_i f(X_1, \dots, X_n) (z - X_i) \right)^2 \right]$$
$$\leq \mathbf{E} \left[ (b - a)^2 \sum_{i=1}^n \nabla_i f(X_1, \dots, X_n)^2 \right]$$
$$= (b - a)^2 \mathbf{E} \left[ \left\| \nabla_i f(X_1, \dots, X_n) \right\|^2 \right].$$

c. Since f is L-Lipschitz, we know that  $||f(x)|| \leq L$  for all  $x \in [a, b]^n$ . Thus, from part b.,

$$\operatorname{Var}[f(X_1,\ldots,X_n)] \le L^2(b-a)^2.$$

## 2.2 Markov semigroups

**Problem 2.6** (Some elementary identities). Let  $P_t$  be a Markov semigroup with generator  $\mathcal{L}$  and stationary measure  $\mu$ . Prove the following elementary facts:

- a) Show that  $\mu(\mathcal{L}f) = 0$  for every  $f \in \text{Dom}(\mathcal{L})$ .
- b) If  $\phi \colon \mathbb{R} \to \mathbb{R}$  is convex, then  $P_t \phi(f) \ge \phi(P_t f)$  when  $f, \phi(f) \in L^2(\mu)$ .
- c) If  $\phi \colon \mathbb{R} \to \mathbb{R}$  is convex, then  $\mathcal{L}\phi(f) \ge \phi'(f)\mathcal{L}f$  when  $f, \phi(f) \in \text{Dom}(\mathcal{L})$ .
- d) Let  $f \in \text{Dom}(\mathcal{L})$ . Show that the following process is a martingale:

$$M_t^f := f(X_t) - \int_0^t \mathcal{L}f(X_s) \mathrm{d}s.$$

Solution. a) Recall that the generator is defined as a limit in  $L^2(\mu)$ , i.e. we have

$$\lim_{t \to 0} \left\| \frac{1}{t} (P_t f - f) - \mathcal{L} f \right\|_{L^2(\mu)} \to 0.$$

In particular, this implies

$$\lim_{t \to 0} \int \frac{1}{t} \left( P_t f - f \right) \mathrm{d}\mu = \int \mathcal{L} f \mathrm{d}\mu.$$

Since stationarity implies for all t

$$\frac{1}{t} \int (P_t f - f) \, \mathrm{d}\mu = \frac{1}{t} \mathbb{E} \left[ \mathbb{E} \left[ f(X_t) \mid X_0 \right] - f(X_0) \right] = 0$$

we can conclude

$$\int \mathcal{L}f d\mu = \lim_{t \to 0} \int \frac{1}{t} \left( P_t f - f \right) d\mu = 0.$$

b) This is just the conditional version of Jensen's inequality, i.e. for any x

$$P_t \phi(f(x)) = \mathbb{E}\left[\phi(f(X_t)) \mid X_0 = x\right] \ge \phi\left(\mathbb{E}\left[f(X_t) \mid X_0 = x\right]\right) = \phi(P_t f(x)).$$

c) By convexity we have for any y, z that  $\phi(y) - \phi(z) \ge \phi'(z)(y-z)$ . Using Jensen's inequality from above and with  $y = P_t f(x)$  and z = f(x), we get

$$\frac{P_t(\phi(f_x)) - \phi(f(x))}{t} \ge \frac{\phi(P_t f(x)) - \phi(f(x))}{t}$$
$$\ge \frac{\phi'(f(x)) \left(P_t f(x) - f(x)\right)}{t}$$

Taking the limit  $t \to 0$  yields the required result.

d) We have

$$M_t^f := f(X_t) - \int_0^t \mathcal{L}f(X_s) \mathrm{d}s.$$

Let  $\{\mathcal{F}_t\}_{t\geq 0}$  denote the natural filtration with respect to  $X = (X_t)$  and  $\tau \leq t$ , then

$$\mathbb{E}\left[M_t^f \mid \mathcal{F}_\tau\right] = \mathbb{E}\left[f(X_t) \mid \mathcal{F}_\tau\right] - \mathbb{E}\left[\int_0^t \mathcal{L}f(X_s) \mathrm{d}s \mid \mathcal{F}_\tau\right]$$
$$= \mathbb{E}\left[f(X_t) \mid X_\tau\right] - \mathbb{E}\left[\int_\tau^t \mathcal{L}f(X_s) \mathrm{d}s \mid X_\tau\right] - \mathbb{E}\left[\int_0^\tau \mathcal{L}f(X_s) \mathrm{d}s \mid \mathcal{F}_\tau\right]$$
$$= P_{t-\tau}f(X_\tau) - \mathbb{E}\left[\int_0^\tau \mathcal{L}f(X_s) \mathrm{d}s \mid X_\tau\right] - \int_0^\tau \mathcal{L}f(X_s) \mathrm{d}s$$
$$= P_{t-\tau}f(X_\tau) - P_{t-\tau}f(X_\tau) + f(X_\tau) - \int_0^\tau \mathcal{L}f(X_s) \mathrm{d}s$$
$$= f(X_\tau) - \int_0^\tau \mathcal{L}f(X_s) \mathrm{d}s = M_\tau^f,$$

where we used that  $\int_0^{\tau} \mathcal{L}f(X_s) ds$  is  $\mathcal{F}_{\tau}$  measurable in the second line and the Markov property in the third.